# Probabilistic Methods in Combinatorics Lecture 8 

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\begin{gathered}
\operatorname{Pr}\left(X_{i}=1\right)=p, \quad \operatorname{Pr}\left(X_{i}=0\right)=1-p \\
B(n, p) \sim N(\mu, \sigma)
\end{gathered}
$$

$N(\mu, \sigma)$ : normal distribution with $\mu=n p$ and $\sigma=\sqrt{n p(1-p)}$.

## Large deviations



## Large deviation inequality I

Chernoff inequalities: Suppose $X=\sum_{i=1}^{n} X_{i}$, where $X_{i}$ are independent 0-1 random variables with

$$
\operatorname{Pr}\left(X_{i}=1\right)=p_{i}, \quad \operatorname{Pr}\left(X_{i}=0\right)=1-p_{i} .
$$

Then we have

$$
\begin{aligned}
& \operatorname{Pr}(X<E(X)-\lambda) \leq e^{-\frac{\lambda^{2}}{2 E(X)}} \\
& \operatorname{Pr}(X>E(X)+\lambda) \leq e^{-\frac{\lambda^{2}}{2(E(X)+\lambda / 3)}}
\end{aligned}
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- $\quad \nu=\sum_{i=1}^{n} a_{i}^{2} p_{i}$


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Theorem [Chung,Lu] We have

$$
\begin{align*}
& \operatorname{Pr}(X<E(X)-\lambda) \leq e^{-\lambda^{2} / 2 \nu}  \tag{1}\\
& \operatorname{Pr}(X>E(X)+\lambda) \leq e^{-\frac{\lambda^{2}}{2(\nu+M \lambda / 3)}} \tag{2}
\end{align*}
$$

## Large deviation inequality III

Theorem [McDiarmid]: Suppose $X_{1}, X_{2}, \ldots, X_{n}$ are independent random variables with $X_{i}-E\left(X_{i}\right) \leq M$ for a positive constant $M$. Let $X=\sum_{i=1}^{n} X_{i}$. Then

$$
\operatorname{Pr}(X-E(X)>\lambda) \leq e^{-\frac{\lambda^{2}}{2(\operatorname{Var}(X)+M \lambda / 3)}} .
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Note: If $\operatorname{Pr}\left(X_{i}=a_{i}\right)=p_{i}$ and $\operatorname{Pr}\left(X_{i}=0\right)=1-p_{i}$, then $\operatorname{Var}(X)=a_{i}^{2} p_{i}\left(1-p_{i}\right) \leq \nu$. Thus

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\operatorname{Pr}(X-E(X)>\lambda) \leq e^{-\frac{\lambda^{2}}{2(\nu+M \lambda / 3)}} .
$$

This theorem implies inequality of upper tail in previous Theorem.

## Large deviation inequality IV

Theorem [Chung, Lu] Suppose $X_{i}$ are independent random variables satisfying $X_{i} \leq M$, for $1 \leq i \leq n$. Let $X=\sum_{i=1}^{n} X_{i}$ and $\|X\|=\sqrt{\sum_{i=1}^{n} \mathrm{E}\left(X_{i}^{2}\right)}$. Then we have

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\operatorname{Pr}(X \geq E(X)+\lambda) \leq e^{-\frac{\lambda^{2}}{2\left(\|X\|^{2}+M \lambda / 3\right)}} .
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$$
\begin{aligned}
& \text { Let } X_{i}^{\prime}=X_{i}-\mathrm{E}\left(X_{i}\right), \text { and } X^{\prime}=X-\mathrm{E}(X) . \\
& \qquad X-\mathrm{E}(X)=X^{\prime}-\mathrm{E}\left(X^{\prime}\right) \\
& \left\|X^{\prime}\right\|^{2}=\sum_{i=1}^{n} \mathrm{E}\left(X_{i}^{\prime 2}\right)=\operatorname{Var}(X) .
\end{aligned}
$$

## Lower tail

Theorem [Chung, Lu] Suppose $X_{i}$ are independent random variables satisfying $X_{i} \geq 0$, for $1 \leq i \leq n$. Let $X=\sum_{i=1}^{n} X_{i}$ and $\|X\|=\sqrt{\sum_{i=1}^{n} \mathrm{E}\left(X_{i}^{2}\right)}$. Then we have

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\operatorname{Pr}(X \leq E(X)-\lambda) \leq e^{-\frac{\lambda^{2}}{2\|X\|^{2}}}
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\operatorname{Pr}(X \leq E(X)-\lambda) \leq e^{-\frac{\lambda^{2}}{2\|X\|^{2}}} .
$$

Proof: Let $X_{i}^{\prime}=-X_{i}$ and $X^{\prime}=-X$. Applying the upper tail to $X^{\prime}$ with $M=0$, we get

$$
\begin{aligned}
\operatorname{Pr}(X \leq E(X)-\lambda) & =\operatorname{Pr}\left(X^{\prime} \geq E\left(X^{\prime}\right)+\lambda\right) \\
& \leq e^{-\frac{\lambda^{2}}{2\left\|X^{\prime}\right\|^{2}}}=e^{-\frac{\lambda^{2}}{2\|X\|^{2}}}
\end{aligned}
$$

## A special function

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g(y)=2 \sum_{k=2}^{\infty} \frac{y^{k-2}}{k!}=\frac{2\left(e^{y}-1-y\right)}{y^{2}} .
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- $g(0)=1$.


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- $g(0)=1$.
- $g(y) \leq 1$, for $y<0$.
- $g(y)$ is monotone increasing, for $y \geq 0$.
- For $y<3$, we have

$$
g(y)=2 \sum_{k=2}^{\infty} \frac{y^{k-2}}{k!} \leq \sum_{k=2}^{\infty} \frac{y^{k-2}}{3^{k-2}}=\frac{1}{1-y / 3} .
$$

## Proof of upper tail

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& =\prod_{i=1}^{n} \mathrm{E}\left(1+t \mathrm{E}\left(X_{i}\right)+\frac{1}{2} t^{2} X_{i}^{2} g\left(t X_{i}\right)\right)
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& \leq \prod_{i=1}^{n}\left(1+t \mathrm{E}\left(X_{i}\right)+\frac{1}{2} t^{2} \mathrm{E}\left(X_{i}^{2}\right) g(t M)\right)
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& \leq \prod_{i=1}^{n} \int_{\mathrm{E}}^{t \mathrm{E}\left(X_{i}\right)+\frac{1}{2} t^{t} \mathrm{E}\left(X_{i}^{2}\right) g(t M)}
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& \leq \prod_{i=1}^{n} e^{t \mathrm{E}\left(X_{i}\right)+\frac{1}{2} t^{2} \mathrm{E}\left(X_{i}^{2}\right) g(t M)} \\
& =e^{t \mathbb{E}(X)+\frac{1}{2} t^{2} g(t M)\|X\|^{2}} .
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## Large deviation inequality V

Theorem [Chung, Lu] Let $X_{i}$ denote independent random variables satisfying $X_{i} \leq \mathrm{E}\left(X_{i}\right)+a_{i}+M$, for $1 \leq i \leq n$. For, $X=\sum_{i=1}^{n} X_{i}$, we have

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- $\quad X_{i}^{\prime} \leq M$ for $1 \leq i \leq n$.


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## continue

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\operatorname{Pr}(X \geq E(X)+\lambda) & =\operatorname{Pr}\left(X^{\prime} \geq E\left(X^{\prime}\right)+\lambda\right) \\
& \leq e^{-\frac{\lambda^{2}}{\left.2\left(\| X^{\prime}\right)^{2}+M \lambda / 3\right)}} \\
& =e^{-\frac{\lambda^{2}\left(\lambda ^ { 2 } \left(\lambda^{2}\right.\right.}{2\left(\operatorname{Tar}(X)+\sum_{i=1}^{n} a_{i}^{2}+M \lambda / 3\right)}} .
\end{aligned}
$$

It remains to verify

$$
\begin{aligned}
& X^{\prime}-\mathrm{E}\left(X^{\prime}\right)=X-\mathrm{E}(X) \\
& \left\|X^{\prime}\right\|^{2}=\operatorname{Var}(X)+\sum_{i=1}^{n} a_{i}^{2}
\end{aligned}
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## Large deviation inequality VI

Theorem [Chung, Lu] Suppose $X_{i}$ are independent random variables satisfying $X_{i} \leq \mathrm{E}\left(X_{i}\right)+M_{i}$, for $0 \leq i \leq n$. We order $X_{i}$ 's so that $M_{i}$ are in an increasing order. Let $X=\sum_{i=1}^{n} X_{i}$. Then for any $1 \leq k \leq n$, we have

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- $M$ is replaced by $M_{k}$.
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- McDiarmid's inequality is a special case with $k=n$.


## Proof

For fixed $k$, we choose $M=M_{k}$ and

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a_{i}= \begin{cases}0 & \text { if } 1 \leq i \leq k \\ M_{i}-M_{k} & \text { if } k \leq i \leq n\end{cases}
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\sum_{i=1}^{n} a_{i}^{2}=\sum_{i=k}^{n}\left(M_{i}-M_{k}\right)^{2} .
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Apply previous theorem with these $a_{i}$ 's.

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## Expectation and Variance

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\operatorname{Var}(X) & =\sum_{i=1}^{n} \operatorname{Var}\left(X_{i}\right) \\
& =(n-1) p(1-p)+n p(1-p) \\
& =(2 n-1) p(1-p)
\end{aligned}
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In particular, for constant $p \in(0,1)$ and $\lambda=\Theta\left(n^{\frac{1}{2}+\epsilon}\right)$, we have

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\operatorname{Pr}(X \geq \mathrm{E}(X)+\lambda) \leq e^{-\Theta\left(n^{\epsilon}\right)}
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